

$$\begin{matrix} a & b \\ c & d \end{matrix}$$

$$ac + bd$$

$$(a+b)(c+d)$$

Let B_w and C_w denote the # of codewords of weight w , in the first $0:n$ code-bits & the last $0:n$ code-bits. Let D_w denote the resultant weight distrib.

$$D_w = \sum_{i=0}^w B_i C_{w-i}$$

$$S = \sum_{w=0}^n D_w P^w = \sum_{w=0}^n P^w \sum_{i=0}^w B_i C_{w-i}$$

B_0	B_1	B_2	...	$B_{0:n-1}$	$B_{0:n}$
C_0	C_1	C_2	...	$C_{0:n-1}$	$C_{0:n}$

C.S. inequality:

$$S \leq \sum_{w=0}^n P^w \sqrt{\sum_{i=0}^w B_i^2} \sqrt{\sum_{i=0}^w C_i^2}$$

$$S = \sum_{w=0}^n \sum_{i=0}^w [(B_i P^i)(C_{w-i} P^{w-i})]$$

$$\leq \sum_{w=0}^n \left[\sum_{i=0}^w B_i P^i \right] \left[\sum_{i=0}^w C_{w-i} P^{w-i} \right], \text{ since } B_i, C_i$$

$$= \sum_{w=0}^n \left[\sum_{i=0}^w B_i P^i \right] \left[\sum_{i=0}^w C_i P^i \right]$$

JUNE '10

DIMAKIS ET. AL. - RD FN. FOR LDGM CODE

Sequence $S \sim \text{iid Bernoulli } p = 1/2$. $S \in \{0, 1\}^n$

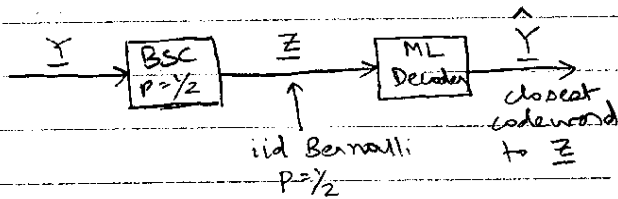
↓ Quantize

$\hat{Z} \in \{0, 1\}^m$

↓ Reconstruct

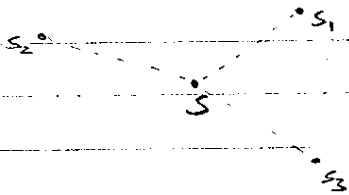
$$\hat{S} = G \hat{Z}, \quad \hat{Z} \in \{0, 1\}^m, \quad G \in \{0, 1\}^{n \times m}$$

Connection with BSC-decoding:



ML-decoding:

Given S , can identify a set $\{\hat{z}_i\}$ of codewords that are closest (at same distance) to S .



If the codeword is unique, can then relate the entropy $H(\hat{Z}|S)$ to the rank of G . Difficulty is in dealing with the case that there are multiple codewords at the nearest neighbour distance.

1/06/10

ANALYZING COMPONENT DUAL CODES SEPARATELY:

$$G = \left[\underbrace{G_1}_{k \times \theta_1 n} \quad \underbrace{G_2}_{k \times \theta_2 n} \right]_{k \times n} \quad \theta_1 + \theta_2 = 1$$

$$q_v = \sum_{u: \langle v|u \rangle = 0} p_u \quad v = \left[\underbrace{v_1}_{1 \times \theta_1 n} \quad \underbrace{v_2}_{1 \times \theta_2 n} \right] \quad u = [u_1; u_2]$$

$$\{u \mid \langle v|u \rangle = 0\} = \left\{ u \mid \begin{array}{l} \langle v_1|u_1 \rangle = 0 \ \& \ \langle v_2|u_2 \rangle = 0 \\ \text{OR} \ \langle v_1|u_1 \rangle = 1 \ \& \ \langle v_2|u_2 \rangle = 1 \end{array} \right\}$$

$$1 - q_v = \underbrace{\sum_{\substack{u: \langle v_1|u_1 \rangle = 0 \\ \langle v_2|u_2 \rangle = 0}} p_u}_{\triangleq q_v^1} + \underbrace{\sum_{\substack{u: \langle v_1|u_1 \rangle = 1 \\ \langle v_2|u_2 \rangle = 1}} p_u}_{\triangleq q_v^2}$$

$$Y = XG$$

$$= X[G_1 \ G_2]$$

$$[Y_1 \ Y_2] = [XG_1 \ XG_2]$$

$$\begin{array}{c} \mathcal{C} \\ \downarrow \\ Z_1, Z_2 \end{array} \quad \downarrow \mathcal{C} = \text{BSC}(p)$$

$$z_i = Y_i + \xi_i$$

$$q_v^z = P_n \left\{ v_1 z_1^T = 1, v_2 z_2^T = 1 \right\}$$

Suppose that v is such that $v_1 \notin C_1^\perp, v_2 \notin C_2^\perp$

$$q_v^z = P_n \left\{ v_1 (Y_1 + \xi_1) = 1, v_2 (Y_2 + \xi_2) = 1 \right\}$$

CHECK!

$$\equiv P_n \left\{ \beta_1 = 1, \beta_2 = 1 \right\}, \quad \beta_1, \beta_2 \sim \text{iid Bernoulli}$$

$$= \frac{1}{4} = q_v^z$$

$$v_1 \in C_1^\perp, v_2 \in C_2^\perp:$$

$$q_v^z = P_n \left\{ \sum_{i: (v_1)_i=1} \xi_i = 1, \sum_{i: (v_2)_i=1} \xi_i = 1 \right\}$$

$$= P_n \left\{ \sum_{i: (v_1)_i=1} \xi_i = 1 \right\} \cdot P_n \left\{ \sum_{i: (v_2)_i=1} \xi_i = 1 \right\}$$

$$= \frac{1 - \prod_{i: (v_1)_i=1} (1-2p_i)}{2} \cdot \frac{1 - \prod_{i: (v_2)_i=1} (1-2p_i)}{2}$$

$$v_1 \in C_1^\perp, v_2 \notin C_2^\perp:$$

$$q_v^z = P_n \left\{ \sum_{i: (v_2)_i=1} \xi_i = 1, \beta_2 = 1 \right\}$$

$$= \frac{1}{2} \cdot \left[\frac{1 - \prod_{i: (v_2)_i=1} (1-2p_i)}{2} \right]$$

$$v_1 \notin C_1^\perp, v_2 \in C_2^\perp: \quad q_v^z = \frac{1}{2} \left[\frac{1 - \prod_{i: (v_2)_i=1} (1-2p_i)}{2} \right]$$

$$v_1 \in C_1^\perp, v_2 \in C_2^\perp \stackrel{\Delta}{=} a \quad \stackrel{\Delta}{=} b$$

$$q_v^1 = \left(\frac{1 + \frac{\prod_{i, (v_i)_i=1} (1-2p_i)}{2}}{2} \right) \left(\frac{1 + \frac{\prod_{i, (v_i)_i=1} (1-2p_i)}{2}}{2} \right)$$

$$v_1 \in C_1^\perp, v_2 \notin C_2^\perp :$$

$$q_v^1 = \frac{1}{2} \left[\frac{1 + \frac{\prod_{i, (v_i)_i=1} (1-2p_i)}{2}}{2} \right]$$

$$v_1 \notin C_1^\perp, v_2 \in C_2^\perp :$$

$$q_v^1 = \frac{1}{2} \left[\frac{1 + \frac{\prod_{i, (v_i)_i=1} (1-2p_i)}{2}}{2} \right]$$

$$v_1 \notin C_1^\perp, v_2 \notin C_2^\perp : 1 - q_v = \frac{1}{2}$$

$$v_1 \in C_1^\perp, v_2 \notin C_2^\perp : 1 - q_v = \frac{1}{2} \left[\frac{1}{2} + \frac{1}{2} \right] = \frac{1}{2}$$

$$v_1 \notin C_1^\perp, v_2 \in C_2^\perp : 1 - q_v = \frac{1}{2}$$

$$v_1 \in C_1^\perp, v_2 \in C_2^\perp :$$

$$+ q_v = \frac{1}{4} \left[(1+a)(1+b) + (1-a)(1-b) \right]$$

$$= \frac{1}{4} \left[1 + \cancel{a} + \cancel{b} + ab + 1 - \cancel{a} - \cancel{b} + ab \right]$$

$$= \frac{1}{2} [1 + ab]$$

$$= \frac{1}{2} \left[1 + \frac{\prod_{i, (v_i)_i=1} (1-2p_i)}{2} \cdot \frac{\prod_{i, (v_i)_i=1} (1-2p_i)}{2} \right]$$

$$= \frac{1}{2} \left[1 + \frac{\prod_{i, v_i=1} (1-2p_i)}{2} \right]$$

$$\Rightarrow q_v = \frac{1}{2} \left[1 - \frac{\prod_{i, v_i=1} (1-2p_i)}{2} \right]$$

DEPENDENCE BETWEEN SUMS OF CODEWORDS:

$$P_{\mathcal{C}} \{ v_1^T (y_1 + \xi_1) = 1, v_2^T (y_2 + \xi_2) = 1 \}$$

$$v_1^T y_1 \neq 0, v_2^T y_2 \neq 0$$

Each component of y_i is a sum of d_i randomly chosen info. symbols.

Let w_1 & w_2 denote the weights of y_1 and y_2 .

Have a total of k info. symbols, iid Bernoulli(p).
Sum 1: Pick $w_1 d_1$ from the info. symbols (with replacement) and add w_1 iid Bernoulli(p) variables to it.
Sum 2: $w_2 d_2$ from info. symbols + w_2 iid Bernoulli(p)

$$\begin{aligned} \text{Need } & P(\text{sum 1} = 1, \text{sum 2} = 1) \\ & \leq P(\text{sum 1} + \text{sum 2} = 0) \end{aligned}$$

If there is one info. symbol that appears solely in sum i and not sum j ($i \neq j$), then sum 1 and sum 2 are independent. What is this probability equal to?

One particular way in which at least one symbol is distinct between the two sums: for sum 1, pick the first variable, and pick the remaining $(w_1 d_1 - 1)$ from the remaining $(k-1)$ variables. For sum 2, pick $w_2 d_2$ variables from the $(k-1)$ variables.

of ways in which this can be done:

$$k \cdot (k-1)^{w_1 d_1 - 1} \cdot (k-1)^{w_2 d_2}$$

Total # of ways of picking the two sums: $k^{w_1 d_1} \cdot k^{w_2 d_2}$

$$\text{Prob}(\text{at least one distinct}) \geq \frac{k \cdot (k-1)^{w_1 d_1 + w_2 d_2}}{k^{w_1 d_1 + w_2 d_2}} = \left(\frac{k-1}{k} \right)^{w_1 d_1 + w_2 d_2 - 1}$$

At worst, w_i grows as $\alpha_i k$

$$\begin{aligned}
 P_n(\text{at least one distinct}) &\geq \left(\frac{k-1}{k}\right)^{k(d_1 \alpha_1 + d_2 \alpha_2) - 1} \\
 &= \left(\frac{k-1}{k}\right)^{\beta k - 1} \quad \text{some const.} \\
 &= \left(1 - \frac{1}{k}\right)^{\beta k} \approx \frac{1}{1 - \frac{1}{k}}
 \end{aligned}$$

$$\left\{ e^x = \lim_{n \rightarrow \infty} \left(1 + \frac{x}{n}\right)^n \right\}$$

$$\lim_{k \rightarrow \infty} P_n \{ \text{at least one distinct} \} \geq e^{-\beta}$$

Suppose that we exclude one variable, ^(info symbol) from the set of codewords with degree d_1 , and only wish it in the set with degree d_2 .

What is the prob. that the $w_2 d_2$ variables picked for sum 2 contains this particular info symbol?

Prob. that it will appear:

$$\frac{1 \cdot (k-1)^{w_2 d_2 - 1} + 1 \cdot (k-1)^{w_2 d_2 - 3} + 1 \cdot (k-1)^{w_2 d_2 - 5} + \dots + (k-1)^0}{k^{w_2 d_2}}$$

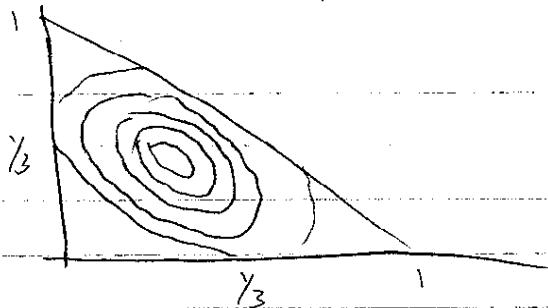
$$= 1 \cdot [(k-1)^2]$$

↑ if $w_2 d_2$ is odd

$$x, y, z \geq 0, \quad x+y+z=1$$

$$H(x, y, z) = \tilde{H}(x, y)$$

$$\tilde{H}(x, y) = \mu$$



$$H(z)$$

$$P_n \xrightarrow{H} q_n$$

$$\|P_n - H_n^{-1}\| = \|q_n - \pi\|$$

8/06/10

1.1.D. RANDOM CASE :

Analyze the quantity $\Pr \{ I(x, z) < n \text{Cap}(C) \}$ for the iid random ensemble. This will allow us to see if our bounding techniques are responsible for the looseness in threshold, or if it is a limitation of the fact that we aren't choosing an appropriate backoff from capacity. See if the random coding case provides some hints on choosing good backoffs.

Look at the bounds of Macmillan & Gallager. Try to use some facts that the coset of a random code is yet another random code, and see if this bound can be analysed.

28/06/10

WEIGHT DISTRIBUTION OF RANDOM LINEAR CODES

Random linear code $C(n, k)$ in \mathbb{F}_q^n

$G \in \mathbb{F}_q^{k \times n}$: entries chosen iid from \mathbb{F}_q with prob.

A_i : # of codewords of Hamming weight i ,

$$i = 0, 1, \dots, n$$

$$x \log \left(\frac{1}{x} \right)$$

$$f(E(x)) \geq E[f(x)]$$

$$W_c(x) = \sum_{i=0}^n A_i x^i$$

$$E[W_c(x)] = 1 + \frac{q^k - 1}{q^n} (1 + (q-1)x)^n$$

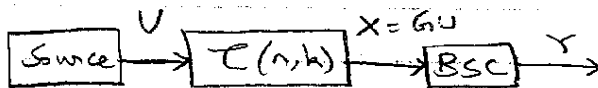
See Blinovskiy, Erez & Litsyn,
"weight distribution moments of random linear
coset codes"
Des. Codes Cryptogr.

Random coset code:

$$\tilde{c}(n, k) = c(n, k) + z_0, \quad z_0 \sim \text{unif}(\mathbb{F}_q^n)$$

$$E[W_{\tilde{c}}(x)] = q^{k-n} [1 + (q-1)x]^n$$

MACMULLAN & COLLINS' RESULT:



It is assumed that X is an invertible function of U , i.e., that G is of rank k .

$$C_{YX} = C_{YU} = k + H(R) - nH(p), \quad \text{where}$$

$$H(R) = \sum_{i=1}^{2^{n-k}} \sum_{j=0}^n w(i, j) p^j (1-p)^{n-j} \log \left[\frac{1}{\sum_{j=0}^n w(i, j) p^j (1-p)^{n-j}} \right]$$

$w(i, j) \rightarrow$ # of n -tuples of Hamming weight j in i th row of the standard array of $C(n, k)$

In general, the rank is not always full when picking all entries of G iid uniformly at random from \mathbb{F}_q . Pdf of the rank known, see Levitzky

Suppose that we have a many to one mapping:

$$U \xrightarrow[\text{one}]{\text{many to}} X \rightarrow Y$$

Suppose that groups of n number of values of U get mapped to a particular X

Claim: $H(Y/X) = H(Y/U)$ (see AWGN writeup of a proof)

Hence the invertibility assumption for $U \rightarrow X$ is not needed in Macmillan-Collins.

Let $W_{c,i} \rightarrow$ weight enumerating fn. for the i th row of the std. array

$$\begin{aligned} H(R) &= \sum_{i=1}^{2^{n-k}} \sum_{j=0}^n w(i,j) P^j (1-P)^{n-j} \log \left[\frac{1}{(1-P)^n W_{c,i} \left(\frac{P}{1-P}\right)} \right] \\ &= \sum_{i=1}^{2^{n-k}} \log \left[\frac{1}{(1-P)^n W_{c,i} \left(\frac{P}{1-P}\right)} \right] (1-P)^n W_{c,i} \left(\frac{P}{1-P}\right) \end{aligned}$$

EVALUATING OUR BOUNDS FOR THE RANDOM ENSEMBLE

$$P_n \{ I(X_i Z) < n \text{Cap}(\epsilon) \} \leq \log_2 \left(\sum_{w=0}^n E[B_w] (1-2\epsilon)^{2w} \right)$$

$$= \log_2 \left(E \{ W_c ((1-2\epsilon)^2)^w \} \right)$$

Incorrect, need to evaluate the dual side (this is what Bw stands for)

$$= \log_2 \left[1 + \frac{2^k - 1}{2^n} \left(1 + (1-2\epsilon)^2 \right)^n \right]$$

$$\left(2^k - 1 \right) \left(\frac{1 + (1-2\epsilon)^2}{2} \right)^n \leq \left(2^R \left(\frac{1 + (1-2\epsilon)^2}{2} \right) \right)^n \rightarrow 0$$

$2^{R-1} [1 + (1-2\epsilon)^2] < 1$

$$2^R < \frac{2}{1 + (1-2p)^2}$$

$$R < \log_2 \left[\frac{2}{1 + (1-2p)^2} \right]$$

Does not make sense. \checkmark
like the condition to be

Wadayama's paper (Arxiv, Mar. 2008):

Random LDPC ensemble has weight distribution

$$\mathbb{E}[B_w] = \binom{n}{w} 2^{-(n-k)} \quad \text{for } n \geq 1, w \geq 1$$

$$\begin{aligned} \sum_{w=0}^n \mathbb{E}[B_w] (1-2p)^{2w} &= 1 + 2^{-(n-k)} \left[\sum_{w=1}^n \binom{n}{w} [(1-2p)^2]^w \right] \\ &= 1 + 2^{-(n-k)} \left\{ [1 + (1-2p)^2]^n - 1 \right\} \end{aligned}$$

$$2^{-(n-k)} \left\{ [1 + (1-2p)^2]^n - 1 \right\} \leq 2^{-(n-k)} [1 + (1-2p)^2]^n$$

CHECK:

$\mathbb{E}[w_c(i)]$ should be equal to q^k

$$\begin{aligned} \text{Blinovsky: } \mathbb{E}[w_c(i)] &= 1 + \frac{q^k - 1}{q^n} \left(1 + (q-1)(1) \right)^n \\ &= q^k \quad \checkmark \end{aligned}$$

$$\text{Wadayama: } 1 + \sum_{w=1}^n \binom{n}{w} 2^{-(n-k)}$$

$$= 1 + 2^{-(n-k)} \cdot (2^n - 1)$$

$$= 1 + 2^k - 2^{n-k} \neq 2^k \quad (??)$$

30/06/10

CALCULATING THE WEIGHT ENUMERATOR FOR THE RANDOM ENSEMBLE:

Along the lines of Blinovsky et al.,

$$W_c(x, y) = \sum_{i=0}^n A_i x^i y^{n-i}$$

$$= \sum_{m=0}^n \left(\sum_{i=0}^{q^k-1} \xi_{im} \right) x^m y^{n-m}$$

$$E(\xi_{im}) = \begin{cases} 1, & i=0, m=0 \\ 0, & i=0, m \neq 0 \\ \frac{1}{q^n} \binom{n}{m} (q-1)^m, & \text{otherwise} \end{cases}$$

$$W_c(x, y) = y^n + \sum_{m=0}^n \sum_{i=1}^{q^k-1} \frac{1}{q^n} \binom{n}{m} (q-1)^m x^m y^{n-m}$$

$$= y^n + \frac{q^k-1}{q^n} \sum_{m=0}^n \binom{n}{m} [x(q-1)]^m y^{n-m}$$

$$= y^n + \frac{q^k-1}{q^n} [y + x(q-1)]^n$$

MacWilliams identity:

$$W_{c^\perp}(x, y) = \frac{1}{|C|} W_c(y-x, y+x)$$

The problem in using MacWilliams identity for the ensemble is that the code size is random as well